

Bi-fidelity Surrogate Modelling: Showcasing the need for new test instances - Appendix

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Appendix A: Kriging

Kriging was developed by Danie Gerhardus Krige empirically to evaluate mineral resources (Krige, 1951); his work was later formalised by Matheron (1963) and has been expanded on by many scientists. Jones (2001) presents a gentle introduction to Kriging and some of its basic uses; the standard derivation can be seen in Sacks et al. (1989) among others. Kriging is a method developed for (single-source) EBB problems, however it can be applied to Bf-EBB problems by working only with the function f_h . Therefore, in this formulation, f_h is simply denoted by f .

The formulation given by Kriging assumes the function samples made so far at locations $\mathbf{x}_1, \dots, \mathbf{x}_n$ are realisations of random normal variables $Y(\mathbf{x}_1), \dots, Y(\mathbf{x}_n)$ with mean μ and variance σ^2 . Further, the errors are correlated based on the distance between variables, that is

$$Corr(Y(\mathbf{x}_i), Y(\mathbf{x}_j)) = \exp \left\{ - \sum_{k=1}^d \theta_k \|\mathbf{x}_i^k - \mathbf{x}_j^k\|^{p_k} \right\}$$

thus, the multivariate random variable $\mathbf{Y} = [Y(\mathbf{x}_1) \dots Y(\mathbf{x}_n)]$ has the distribution $\mathbf{Y} \sim N(\mathbf{1}\mu, \sigma^2 R)$, with $R_{i,j} = \exp \left\{ - \sum_{k=1}^d \theta_k \|\mathbf{x}_i^k - \mathbf{x}_j^k\|^{p_k} \right\}$. Note this has the hyperparameters $\mu, \sigma^2, \theta_1, \dots, \theta_d, p_1, \dots, p_d$. The values θ_k and p_k give an indication of the effect of moving along any of the dimensions (i.e changing the value of a single variable). The constant θ_k represents how the correlation changes with distance: small values mean there is no correlation even for close points in the k^{th} dimension, but large values indicate even relatively distant sample points (in the k^{th} dimension) are correlated. The constant p_k allows the technique to model from smooth functions ($p_k = 2$) to rough, non-differentiable ones ($p_k \rightarrow 0$).

In order to fit the model to the data, the log density function of \mathbf{Y} is maximised, which after simplification leads to the auxiliary optimisation problem

$$\max_{\theta_1, \dots, \theta_d, p_1, \dots, p_d} -\frac{n}{2} \log(\hat{\sigma}^2) - \frac{1}{2} \log(|R|)$$

with

$$\begin{aligned} \hat{\mu} &= \frac{\mathbf{1}^T R^{-1} \mathbf{y}}{\mathbf{1}^T R^{-1} \mathbf{1}} \\ \hat{\sigma}^2 &= \frac{(\mathbf{y} - \mathbf{1}\hat{\mu})^T R^{-1} (\mathbf{y} - \mathbf{1}\hat{\mu})}{n} \\ \mathbf{y} &= [f(\mathbf{x}_1), \dots, f(\mathbf{x}_n)] \end{aligned}$$

Note this problem cannot be solved analytically and thus the tuning of these hyperparameters is an auxiliary problem that must be solved. Once the model has been trained, for a given sample point \mathbf{x} Kriging provides the most likely objective function $s(\mathbf{x})$ and the variance of the estimate $v^2(\mathbf{x})$. These are given by

$$s(\mathbf{x}) = \hat{\mu} + \mathbf{r}^T R^{-1}(\mathbf{y} - \mathbf{1}\hat{\mu})$$

$$v^2(\mathbf{x}) = \hat{\sigma}^2 \left[1 - \mathbf{r}^T R^{-1} \mathbf{r} + \frac{(1 - \mathbf{1}^T R^{-1} \mathbf{r})^2}{\mathbf{1}^T R^{-1} \mathbf{1}} \right]$$

where

$$\mathbf{r} = \begin{bmatrix} \text{Corr}(Y(\mathbf{x}), Y(\mathbf{x}_1)) \\ \vdots \\ \text{Corr}(Y(\mathbf{x}), Y(\mathbf{x}_n)) \end{bmatrix}$$

Thus, Kriging thinks of the objective function predictor as being a realisation of $Y(\mathbf{x}) \sim N(s(\mathbf{x}), v^2(\mathbf{x}))$. This allows the asking of questions such as what is the uncertainty or what is the expected improvement at a sample point, among others.

Appendix B: Co-Kriging

Kennedy and O'Hagan (2000) present a technique to use multiple information sources combined with Gaussian processes for global optimisation. This technique can be adapted to Kriging (Forrester et al., 2007), producing a new technique known as Co-Kriging. Similarly to Kriging, the idea behind Co-Kriging is to model the responses of the cheap objective function f_l at sample points $\mathbf{X}_l = (\mathbf{x}_1^l, \mathbf{x}_2^l, \dots, \mathbf{x}_{n_l}^l)$ and the responses of the expensive objective function f_h at sample points $\mathbf{X}_h = (\mathbf{x}_1^h, \mathbf{x}_2^h, \dots, \mathbf{x}_{n_h}^h)$ as the realisation of a multivariate random variable:

$$\mathbf{Y} = (\mathbf{Y}_l(\mathbf{X}_l), \mathbf{Y}_h(\mathbf{X}_h)) = (Y_l(\mathbf{x}_1^l), \dots, Y_l(\mathbf{x}_{n_l}^l), Y_h(\mathbf{x}_1^h), \dots, Y_h(\mathbf{x}_{n_h}^h))$$

The multivariate random variable $\mathbf{Y}_l(\mathbf{X}_l)$, that is the response of the cheap objective function, is treated as a multivariate normal random variable with distribution $N(\mu_l, \sigma_l^2 R_l)$. On the other hand, the multivariate random variable $\mathbf{Y}_h(\mathbf{X}_h)$, that is the response of the expensive objective function, is represented by a scaling of ρ of the response of the cheap expensive function $\mathbf{Y}_l(\mathbf{X}_l)$ plus a new Gaussian process \mathbf{Y}_b which models the difference between the cheap and expensive objective functions, that is

$$\mathbf{Y}_h(\mathbf{X}_h) = \rho \mathbf{Y}_l(\mathbf{X}_h) + \mathbf{Y}_b(\mathbf{X}_h)$$

The random variable \mathbf{Y}_b is also treated as a multivariate normal random variable with distribution $N(\mu_b, \sigma_b^2 R_b)$. It is assumed that \mathbf{Y}_l and \mathbf{Y}_b are independent. Thus the following correlation measures are given for \mathbf{Y}_l and \mathbf{Y}_h :

$$\begin{aligned}
Corr(\mathbf{Y}_l(\mathbf{X}_l), \mathbf{Y}_l(\mathbf{X}_l)) &= \sigma_l^2 R_l(\mathbf{X}_l, \mathbf{X}_l) \\
Corr(\mathbf{Y}_h(\mathbf{X}_h), \mathbf{Y}_l(\mathbf{X}_l)) &= \rho \sigma_l^2 R_l(\mathbf{X}_h, \mathbf{X}_l) \\
Corr(\mathbf{Y}_h(\mathbf{X}_h), \mathbf{Y}_h(\mathbf{X}_h)) &= \rho^2 \sigma_l^2 R_l(\mathbf{X}_h, \mathbf{X}_h) + \sigma_b^2 R_b(\mathbf{X}_h, \mathbf{X}_h)
\end{aligned}$$

where

$$\begin{aligned}
R_l(\mathbf{X}_l, \mathbf{X}_l)_{i,j} &= \exp \left\{ - \sum_{k=1}^d \theta_k^l \|(\mathbf{x}_i^l)_k - (\mathbf{x}_j^l)_k\|^{p_k^l} \right\} & 1 \leq i, j \leq n_l \\
R_l(\mathbf{X}_h, \mathbf{X}_l)_{i,j} &= \exp \left\{ - \sum_{k=1}^d \theta_k^l \|(\mathbf{x}_i^h)_k - (\mathbf{x}_j^l)_k\|^{p_k^l} \right\} & 1 \leq i \leq n_h \quad 1 \leq j \leq n_l \\
R_l(\mathbf{X}_h, \mathbf{X}_h)_{i,j} &= \exp \left\{ - \sum_{k=1}^d \theta_k^l \|(\mathbf{x}_i^h)_k - (\mathbf{x}_j^h)_k\|^{p_k^l} \right\} & 1 \leq i, j \leq n_h \\
R_b(\mathbf{X}_h, \mathbf{X}_h)_{i,j} &= \exp \left\{ - \sum_{k=1}^d \theta_k^b \|(\mathbf{x}_i^h)_k - (\mathbf{x}_j^h)_k\|^{p_k^b} \right\} & 1 \leq i, j \leq n_h
\end{aligned}$$

In order to fit the model to the data, the log density function of \mathbf{Y}_l is maximised, which after simplification (Forrester et al., 2007) leads to the auxiliary optimisation problem

$$\max_{\mu_l, \sigma_l^2, \theta_1^l, \dots, \theta_d^l, p_1^l, \dots, p_d^l} - \frac{n_l}{2} \log(\hat{\sigma}_l^2) - \frac{1}{2} \log(|\det(R_l(\mathbf{X}_l, \mathbf{X}_l))|)$$

where

$$\begin{aligned}
\hat{\mu}_l &= \frac{\mathbf{1}^T R_l(\mathbf{X}_l, \mathbf{X}_l)^{-1} \mathbf{Y}_l}{\mathbf{1}^T R_l(\mathbf{X}_l, \mathbf{X}_l)^{-1} \mathbf{1}} \\
\hat{\sigma}_l^2 &= \frac{(\mathbf{Y}_l - \mathbf{1} \hat{\mu}_l)^T R_l(\mathbf{X}_l, \mathbf{X}_l)^{-1} (\mathbf{Y}_l - \mathbf{1} \hat{\mu}_l)}{n_l} \\
\mathbf{Y}_l &= (f_l(\mathbf{x}_1^l), \dots, f_l(\mathbf{x}_{n_l}^l))
\end{aligned}$$

As is the case with Kriging, this problem cannot be solved analytically and is an auxiliary optimisation problem which must be solved. In order to calculate the parameters associated with \mathbf{Y}_b , first \mathbf{b} is defined:

$$\mathbf{b} = \mathbf{Y}_h - \rho \mathbf{Y}_l(\mathbf{X}_h)$$

where $\mathbf{Y}_h = (f_h(\mathbf{x}_1^h), \dots, f_h(\mathbf{x}_{n_h}^h))$, and $\mathbf{Y}_l(\mathbf{X}_h)_i$ is $f_l(\mathbf{x}_i^h)$ if the point has already been evaluated, and otherwise it is $\hat{y}(\mathbf{x}_i^h) = \hat{\mu}_l + \mathbf{r}_l^T R_l(\mathbf{X}_l, \mathbf{X}_l)^{-1} (\mathbf{Y}_l - \mathbf{1} \hat{\mu}_l)$, with $\mathbf{r}_l = (R_l(\mathbf{x}, \mathbf{x}_1^l), \dots, R_l(\mathbf{x}, \mathbf{x}_{n_l}^l))$. That is, if a point has not been evaluated by f_l yet, its Kriging predictor of the cheap model is used instead. A second auxiliary problem is solved to find a second set of hyperparameters, using the log density function of \mathbf{Y}_b :

$$\max_{\rho, \theta_1^b, \dots, \theta_d^b, p_1^b, \dots, p_d^b} -\frac{n_h}{2} \log(\hat{\sigma}_b^2) - \frac{1}{2} \log(|\det(R_b(\mathbf{X}_h, \mathbf{X}_h))|)$$

where

$$\begin{aligned}\hat{\mu}_b &= \frac{\mathbf{1}^T R_b(\mathbf{X}_h, \mathbf{X}_h)^{-1} \mathbf{b}}{\mathbf{1}^T R_b(\mathbf{X}_h, \mathbf{X}_h)^{-1} \mathbf{1}} \\ \hat{\sigma}_b^2 &= \frac{(\mathbf{b} - \mathbf{1}\hat{\mu}_b)^T R_b(\mathbf{X}_h, \mathbf{X}_h)^{-1} (\mathbf{b} - \mathbf{1}\hat{\mu}_b)}{n_h}\end{aligned}$$

Finally, the Co-Kriging predictor is given by

$$s_h(\mathbf{x}) = \hat{\mu} + \mathbf{c}^T C^{-1}(\mathbf{y} - \mathbf{1}\hat{\mu})$$

with

$$C = \begin{bmatrix} \hat{\sigma}_l^2 R_l(\mathbf{X}_l, \mathbf{X}_l) & \hat{\rho} \hat{\sigma}_l^2 R_l(\mathbf{X}_l, \mathbf{X}_h) \\ \hat{\rho} \hat{\sigma}_l^2 R_l(\mathbf{X}_h, \mathbf{X}_l) & \hat{\rho}^2 \hat{\sigma}_l^2 R_l(\mathbf{X}_h, \mathbf{X}_h) + \hat{\sigma}_b^2 R_b(\mathbf{X}_h, \mathbf{X}_h) \end{bmatrix}$$

$$\mathbf{c} = \begin{bmatrix} \hat{\rho} \hat{\sigma}_l^2 R_l(\mathbf{X}_l, \mathbf{x}) \\ \hat{\rho}^2 \hat{\sigma}_l^2 R_l(\mathbf{X}_h, \mathbf{x}) + \hat{\sigma}_b^2 R_b(\mathbf{X}_h, \mathbf{x}) \end{bmatrix}$$

$$\hat{\mu} = \frac{\mathbf{1}^T C^{-1} \mathbf{y}}{\mathbf{1}^T C^{-1} \mathbf{1}}$$

An estimated mean-square error can also be extracted for the predictor, which is given by

$$v^2(\mathbf{x}) = \hat{\rho}^2 \hat{\sigma}_l^2 + \hat{\sigma}_b^2 - \mathbf{c}^T C^{-1} \mathbf{c}$$

The overall algorithm presented by Forrester consists of creating a large set of sample points for the cheap objective function, and then choose a subset of those points to sample the expensive objective function. It then chooses the next sample point by treating the value at a particular point as the realisation of a normal random variable $\sim N(s_h(\mathbf{x}), v^2(\mathbf{x}))$. The chosen sample point is evaluated both by the cheap and expensive objective functions, the models are fitted again, and the next sample point is chosen.

References

- Forrester, A. I., Sóbester, A., and Keane, A. J. (2007). Multi-fidelity optimization via surrogate modelling. *Proceedings of the royal society a: mathematical, physical and engineering sciences*, 463(2088):3251–3269.
- Jones, D. R. (2001). A taxonomy of global optimization methods based on response surfaces. *Journal of global optimization*, 21(4):345–383.
- Kennedy, M. C. and O’Hagan, A. (2000). Predicting the output from a complex computer code when fast approximations are available. *Biometrika*, 87(1):1–13.

- Krige, D. G. (1951). A statistical approach to some basic mine valuation problems on the witwatersrand. *Journal of the Southern African Institute of Mining and Metallurgy*, 52(6):119–139.
- Matheron, G. (1963). Principles of geostatistics. *Economic geology*, 58(8):1246–1266.
- Sacks, J., Welch, W. J., Mitchell, T. J., and Wynn, H. P. (1989). Design and analysis of computer experiments. *Statistical science*, pages 409–423.