

# Online Supplement for “Understanding How Generation Flexibility and Renewable Energy Affect Power Market Competition”

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**Proof of Lemma 1.** The inverse supply function is defined as  $S_k^{-1}(q) = \inf\{p : S_k(p) > q\}$ . Under Assumption 2(i) and (ii),  $S_k^{-1}(0) = p_k^{\min}$  and  $S_k^{-1}(q)$  is continuous and increasing in  $q$  for  $q \geq 0$ . Assumption 2(iii) implies that  $S_k^{-1}(q) > 0$  for  $q > 0$ . Hence,  $C_k(q) = \int_0^q S_k^{-1}(x)dx$  is continuously differentiable, convex, and strictly increasing in  $q$  for  $q \geq 0$ .

Consider the optimal allocation problem in (2), rewritten below

$$C^I(q) = \min \left\{ \sum_{i \in G^I} C_i(q_i) : q_i \geq 0, \sum_{i \in G^I} q_i = q \right\}. \quad (\text{A.1})$$

For any  $\bar{q} > 0$ , the objective in (A.1) is convex on a closed convex set  $\{(q, q_i, i \in G^I) : q \in [0, \bar{q}], q_i \in [0, \bar{q}], \sum_{i \in G^I} q_i = q\}$ . Hence, the theorem on convexity preservation (Heyman and Sobel 1984, p. 525) implies that  $C^I(q)$  is convex in  $q$ .

For a given  $q > 0$ , let the minimizer for (A.1) be  $\{q_i^*\}$ , which has two properties: 1) If  $q_i^* > 0$ , then  $C'_i(q_i^*) = p_0$ , for some constant  $p_0 > 0$  that depends on the given  $q$ ; 2) If  $q_i^* = 0$ , then  $C'_i(0) \geq p_0$ .<sup>1</sup>

Define  $G_+^I \stackrel{\text{def}}{=} \{i \in G^I : C'_i(q_i^*) = p_0\}$ . For  $k \notin G_+^I$ , we have  $q_k^* = 0$  and  $C'_k(0) > p_0$ . Then, for sufficiently small  $\varepsilon > 0$ , we have  $C^I(q + \varepsilon) = \sum_{i \in G_+^I} C_i(q_i^* + \varepsilon_i)$ , for some  $\varepsilon_i \geq 0$  and  $\sum_{i \in G_+^I} \varepsilon_i = \varepsilon$ . Using Taylor's expansion, we can write

$$C^I(q + \varepsilon) = \sum_{i \in G_+^I} [C_i(q_i^*) + \varepsilon_i C'_i(q_i^*) + o(\varepsilon_i)] = C^I(q) + \varepsilon p_0 + o(\varepsilon).$$

Similarly, we can show that  $C^I(q) - C^I(q - \varepsilon) = \varepsilon p_0 + o(\varepsilon)$ . Hence,  $C^I(q)$  is differentiable with derivative  $(C^I)'(q) = p_0 > 0$ .

For  $i \in G_+^I$ , we have  $p_0 = C'_i(q_i^*) = S_i^{-1}(q_i^*)$ , implying  $S_i(p_0) = q_i^*$ . For  $k \notin G_+^I$ , we have  $C'_k(0) > p_0$ , implying  $S_k^{-1}(0) = p_k^{\min} > p_0$ , which in turn implies  $S_k(p_0) = 0 = q_k^*$  due to Assumption 2(i). Hence,  $S^I(p_0) = \sum_{i \in G^I} S_i(p_0) = \sum_{i \in G^I} q_i^* = q$ , which leads to  $p_0 = (S^I)^{-1}(q) = (C^I)'(q)$ . Note that  $(S^I)^{-1}(q)$  is continuous in  $q$  because  $S^I(p)$  satisfies Assumption 2. Hence,  $(C^I)'(q)$  is continuous in  $q$ .

Similar results can be shown for the problem in (3), which completes the proof. ■

**Proof of Theorem 1.** We first prove that (12) is optimal in the scenario of  $L_s - q^I - W_s \geq 0$ . In this scenario, constraints (10)-(11) imply that  $q_s^F \geq L_s - q^I - W_s \geq 0$ . If we set  $q_s^F$  at the lower bound  $L_s - q^I - W_s$ , then  $q_s^V = W_s$  and  $e_s = 0$ , which clearly minimize the objective function in (9).

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<sup>1</sup>To see property 1), note that if  $q_i^*, q_k^* > 0$  and  $C'_i(q_i^*) < C'_k(q_k^*)$ , we can strictly reduce the objective by increasing  $q_i^*$  by  $\varepsilon$  and reducing  $q_k^*$  by  $\varepsilon$ , where  $\varepsilon > 0$  is small. Thus, for  $q_i^*, q_k^* > 0$ ,  $C'_i(q_i^*) = C'_k(q_k^*) = p_0$ . Note that  $p_0 > 0$  because  $C_i(q_i)$  is convex and strictly increasing in  $q_i$ . To see property 2), note that if  $C'_i(0) < p_0 = C'_k(q_k^*)$ , we can strictly reduce the objective by setting  $q_i^* = \varepsilon$  and reducing  $q_k^*$  by  $\varepsilon$ , where  $\varepsilon > 0$  is small.

When  $L_s - q^I - W_s < 0$ , we have  $q_s^{F*} = 0$  because: (i) if  $q_s^F > 0$  and  $e_s > 0$ , then a lower  $q_s^F$  reduces the objective value in (9); (ii) if  $q_s^F > 0$  and  $e_s = 0$ , then  $q_s^V = L_s - q^I - q_s^F < W_s$ , and we can reduce  $q_s^F$  and increase  $q_s^V$  to lower the objective in (9). Hence,  $q_s^{F*} = 0$ , and (9)-(11) becomes

$$\min_{q_s^V \in [0, W_s]} \{ -r q_s^V + h(q^I + q_s^V - L_s) \},$$

where we set  $h(e) = 0$  for  $e < 0$  without affecting the optimal solution. An interior optimal solution satisfies  $h'(q^I + q_s^{V*} - L_s) = r$ , or  $q_s^{V*} = L_s - q^I + \mu(r)$ , which is in the interior if  $0 < L_s - q^I + \mu(r) < W_s$ . If  $L_s - q^I + \mu(r) \geq W_s$ , then  $q_s^{V*} = W_s$ . If  $L_s - q^I + \mu(r) < 0$ , then  $q_s^{V*} = 0$ . This proves (12).

The objective function in (9) is convex on a closed convex set  $\{(q^I, L_s, W_s, q_s^F, q_s^V) : q^I \in [0, \bar{q}], L_s \in [0, \bar{L}], W_s \in [0, \bar{W}], q_s^F \in [0, \bar{q}], (10) \text{ and } (11)\}$  for any  $\bar{q} > 0$ ,  $\bar{L} > 0$ , and  $\bar{W} > 0$ . By the theorem on convexity preservation under minimization (Heyman and Sobel 1984, p. 525) we conclude that  $\widehat{C}(q^I, L_s, W_s)$  is jointly convex in  $(q^I, L_s, W_s)$ .  $\blacksquare$

**Proof of Lemma 2.** We verify the consistency between (13) and (14) by considering each of the four regions. The inequality in (14) is

$$S^F(p) + W \mathbf{1}_{\{p \geq -r\}} - \mu(-p) \geq L - q^I. \quad (\text{A.2})$$

In region  $A_1$ ,  $L - q^I \leq -\mu(r)$ , and (A.2) clearly holds if  $p = p^o \stackrel{\text{def}}{=} -h'(q^I - L)$ . Note that  $p^o \leq -r$ . Thus, for any other price  $p_1 < p^o$ , the left side of (A.2) becomes  $-\mu(-p_1)$ , which is strictly less than  $L - q^I$ . Hence,  $p^o$  is the minimum price for (A.2) to hold.

In region  $A_2$ ,  $L - q^I \in (-\mu(r), W - \mu(r))$ . If  $p = -r$ , then (A.2) holds because  $W - \mu(r) > L - q^I$ . For any other  $p_1 < -r$ , (A.2) does not hold because  $-\mu(-p_1) < -\mu(r) < L - q^I$ .

In region  $A_3$ ,  $L - q^I \in [W - \mu(r), W]$ . If  $p = -h'(q^I + W - L) \in [-r, 0]$ , then (A.2) holds with equality:  $W - (q^I + W - L) = L - q^I$ .

Lastly, in region  $A_4$ ,  $L - q^I > W$ . If  $p = (C^F)'(L - W - q^I) > (C^F)'(0)$ , then (A.2) also holds with equality:  $(L - W - q^I) + W = L - q^I$ .

Hence, the minimum price  $p$  for (A.2) to hold is exactly  $P(q^I, L, W, S^F)$  defined in (13).  $\blacksquare$

**Proof of Theorem 2.** Using Theorem 1, we can write

$$\widehat{C}(q^I, L, W) = C^F(q^{F*}) - r q^{V*} + h(q^I + q^{F*} + q^{V*} - L),$$

where  $q^{F*}$  and  $q^{V*}$  are given in Figure 1 under the four events with the following indicator functions:

$$\begin{aligned} \mathbf{1}_{A_1} &= \mathbf{1}_{L \leq q^I - \mu(r)}, \\ \mathbf{1}_{A_2} &= -\mathbf{1}_{L \leq q^I - \mu(r)} + \mathbf{1}_{L < q^I + W - \mu(r)}, \\ \mathbf{1}_{A_3} &= -\mathbf{1}_{L < q^I + W - \mu(r)} + \mathbf{1}_{L \leq q^I + W}, \\ \mathbf{1}_{A_4} &= \mathbf{1}_{L > q^I + W}. \end{aligned} \quad (\text{A.3})$$

We denote  $\widehat{C}_{A_i}(\cdot, \cdot, \cdot) = \widehat{C}(\cdot, \cdot, \cdot)$  when  $A_i$  occurs. Then, using the optimal policy in Figure 1, we have

$$\begin{aligned}
\widehat{C}_{A_1}(q^I, L, W) &= h(q^I - L), \\
\widehat{C}_{A_2}(q^I, L, W) &= -r(L - q^I + \mu(r)) + h(\mu(r)), \\
\widehat{C}_{A_3}(q^I, L, W) &= -rW + h(q^I + W - L), \\
\widehat{C}_{A_4}(q^I, L, W) &= C^F(L - q^I - W) - rW.
\end{aligned} \tag{A.4}$$

Using (A.3), we can write the expected second-stage cost as

$$\begin{aligned}
\mathbb{E}[\widehat{C}(q^I, L, W)] &= \sum_{i=1}^4 \mathbb{E}[\widehat{C}_{A_i}(q^I, L, W) \mathbf{1}_{A_i}] \\
&= \mathbb{E}\left[(\widehat{C}_{A_1}(q^I, L, W) - \widehat{C}_{A_2}(q^I, L, W)) \mathbf{1}_{L \leq q^I - \mu(r)}\right] \\
&\quad + \mathbb{E}\left[(\widehat{C}_{A_2}(q^I, L, W) - \widehat{C}_{A_3}(q^I, L, W)) \mathbf{1}_{L < q^I + W - \mu(r)}\right] \\
&\quad + \mathbb{E}\left[\widehat{C}_{A_3}(q^I, L, W) \mathbf{1}_{L \leq q^I + W}\right] - \mathbb{E}\left[\widehat{C}_{A_4}(q^I, L, W) \mathbf{1}_{L > q^I + W}\right].
\end{aligned} \tag{A.5}$$

It can be verified that  $\widehat{C}(q^I, L_s, W_s)$  is differentiable in  $q^I$  except at  $q^I = L_s - W_s$ , where the left derivative is  $-(C^F)'(0)$  and the right derivative is  $h'(0)$ . Because  $L$  and  $W$  have continuous distributions,  $\mathbb{E}[\widehat{C}(q^I, L, W)]$  is differentiable in  $q^I$  everywhere. Next, we compute its derivative.

The first three expectations in (A.5) all assume the form of  $\mathbb{E}[g(q^I, L, W) \mathbf{1}_{L \leq b(q^I, W)}]$ , for some functions  $g(q^I, L, W)$  and  $b(q^I, W)$ . Let  $L \in (\underline{L}, \overline{L})$ ,  $W \in (\underline{W}, \overline{W})$ , and let the joint probability density function of  $(L, W)$  be  $f(l, w)$ . We use “ $\vee$ ” and “ $\wedge$ ” for max and min operations. We have

$$\begin{aligned}
\frac{d}{dq^I} \mathbb{E}[g(q^I, L, W) \mathbf{1}_{L \leq b(q^I, W)}] &= \frac{d}{dq^I} \left[ \int_{\underline{W}}^{\overline{W}} \int_{\underline{L}}^{L \vee b(q^I, w) \wedge \overline{L}} g(q^I, l, w) f(l, w) dl dw \right] \\
&= \mathbb{E} \left[ \frac{\partial g(q^I, L, W)}{\partial q^I} \mathbf{1}_{L \leq b(q^I, W)} \right] + \int_{\underline{W}}^{\overline{W}} g(q^I, b(q^I, w), w) f(b(q^I, w), w) \frac{\partial b(q^I, w)}{\partial q^I} \mathbf{1}_{b(q^I, w) \in [\underline{L}, \overline{L}]} dw.
\end{aligned} \tag{A.6}$$

The last expectation in (A.5) is in the form of  $\mathbb{E}[g(q^I, L, W) \mathbf{1}_{L > b(q^I, W)}]$  and its derivative is

$$\begin{aligned}
\frac{d}{dq^I} \mathbb{E}[g(q^I, L, W) \mathbf{1}_{L > b(q^I, W)}] &= \frac{d}{dq^I} \left[ \int_{\underline{W}}^{\overline{W}} \int_{L \vee b(q^I, w) \wedge \overline{L}}^{\overline{L}} g(q^I, l, w) f(l, w) dl dw \right] \\
&= \mathbb{E} \left[ \frac{\partial g(q^I, L, W)}{\partial q^I} \mathbf{1}_{L > b(q^I, W)} \right] - \int_{\underline{W}}^{\overline{W}} g(q^I, b(q^I, w), w) f(b(q^I, w), w) \frac{\partial b(q^I, w)}{\partial q^I} \mathbf{1}_{b(q^I, w) \in [\underline{L}, \overline{L}]} dw.
\end{aligned} \tag{A.7}$$

Now, applying (A.6)-(A.7) to the derivatives of the expectations in (A.5), we find that the integral term  $g(q^I, b(q^I, w), w)$  in (A.6)-(A.7) becomes

$$\begin{aligned}
(\widehat{C}_{A_1}(q^I, l, w) - \widehat{C}_{A_2}(q^I, l, w)) \Big|_{l=q^I - \mu(r)} &= 0, \\
(\widehat{C}_{A_2}(q^I, l, w) - \widehat{C}_{A_3}(q^I, l, w)) \Big|_{l=q^I + w - \mu(r)} &= 0, \\
\widehat{C}_{A_3}(q^I, l, w) \Big|_{l=q^I + w} &= -r w, \\
\widehat{C}_{A_4}(q^I, l, w) \Big|_{l=q^I + w} &= -r w,
\end{aligned}$$

where we have used the value of  $\widehat{C}_{A_i}(\cdot, \cdot, \cdot)$  given by (A.4). This leads to

$$\begin{aligned} \frac{d}{dq^I} \mathbb{E}[\widehat{C}(q^I, L, W)] &= \sum_{i=1}^4 \mathbb{E} \left[ \frac{\partial \widehat{C}_{A_i}(q^I, L, W)}{\partial q^I} \mathbf{1}_{A_i} \right] \\ &= \mathbb{E} \left[ -(C^F)'(L - W - q^I) \mathbf{1}_{A_4} + h'(q^I + W - L) \mathbf{1}_{A_3} + r \mathbf{1}_{A_2} + h'(q^I - L) \mathbf{1}_{A_1} \right] \\ &= \mathbb{E}[-P(q^I, L, W, S^F)], \end{aligned}$$

where the last equality follows from Lemma 1 and the definition in (13).

Now, because the objective in (15) is convex in  $q^I$ ,  $q^{I*} > 0$  is optimal if and only if

$$(C^I)'(q^{I*}) + \frac{d}{dq^I} \mathbb{E}[\widehat{C}(q^{I*}, L, W)] = (C^I)'(q^{I*}) - \mathbb{E}[P(q^{I*}, L, W, S^F)] = 0,$$

which is equivalent to  $(S^I)^{-1}(q^{I*}) = \overline{P}(q^{I*}, S^F)$  or  $q^{I*} = S^I(\overline{P}(q^{I*}, S^F))$ .

On the other hand,  $q^{I*} = 0$  is optimal if and only if  $(C^I)'(0) \geq \overline{P}(0, S^F)$ , which is equivalent to  $S^I(\overline{P}(0, S^F)) = 0$ . (To see the equivalence, note that  $S^I(p) = 0$  if and only if  $p \leq (C^I)'(0)$ .) Thus, the same condition  $q^{I*} = S^I(\overline{P}(q^{I*}, S^F))$  is necessary and sufficient for  $q^{I*} = 0$ . ■

**Proof of Theorem 3. IG's profit function:** Given a first-stage price  $p_0 > 0$ , IG  $i \in G^I$  commits to produce  $S_i(p_0) = \beta_i p_0$ , incurring a cost of  $\frac{1}{2} c_i (\beta_i p_0)^2$ . Thus, the profit is  $\beta_i p_0^2 - \frac{1}{2} c_i (\beta_i p_0)^2 = \beta_i (1 - \frac{1}{2} c_i \beta_i) p_0^2$ . Note that the first-stage price  $p_0$  is equal to the expected second-stage price, which can be written as (using (23)):

$$\overline{P}(q^I, \beta^F) = \frac{q^I}{\beta^I} = \frac{q^{I*}(\beta^I, \beta^F)}{\beta^I}.$$

We define  $\beta_{-i} \stackrel{\text{def}}{=} \beta^I - \beta_i$  and write generator  $i$ 's expected profit as

$$\pi_i(\beta_i; \beta_{-i}, \beta^F) = \frac{\beta_i (1 - \frac{1}{2} c_i \beta_i)}{(\beta_i + \beta_{-i})^2} (q^{I*}(\beta_i + \beta_{-i}, \beta^F))^2, \quad (\text{A.8})$$

which proves (25).

**FG's profit function:** Given the first-stage price  $p_0$ , FG  $j \in G^F$  plans to produce  $S_j(p_0) = \beta_j p_0$  and receives a predispatch payment of  $\beta_j p_0^2$ . When uncertainties are realized and the price is  $p_s$ , generator  $j$ 's actual production is  $\beta_j p_s^+$ , incurring a cost of  $\frac{1}{2} c_j (\beta_j p_s^+)^2$ . The production deviation  $\beta_j (p_s^+ - p_0)$  is settled at  $p_s$ , resulting in a revenue of  $\beta_j (p_s^+ - p_0) p_s$  (can be negative if  $p_s \in (0, p_0)$ ). Thus, the total expected revenue of generator  $j$  is

$$\begin{aligned} \beta_j p_0^2 + \beta_j \mathbb{E}[(p_s^+ - p_0) p_s] &= \beta_j p_0^2 - \beta_j p_0 \mathbb{E}[p_s] + \beta_j \mathbb{E}[p_s^+ p_s] \\ &= \beta_j p_0^2 - \beta_j p_0^2 + \beta_j \mathbb{E}[p_s^+ (p_s^+ + p_s^-)] \\ &= \beta_j \mathbb{E}[(p_s^+)^2], \end{aligned}$$

where the second equality is due to  $p_0 = \mathbb{E}[p_s]$ , which follows from Theorem 2, and the last equality is

due to the fact that  $p_s^+ \cdot p_s^- \equiv 0$ . Thus, the expected profit is

$$\beta_j \mathbf{E}[(p_s^+)^2] - \frac{1}{2} c_j \mathbf{E}[(\beta_j p_s^+)^2] = \beta_j (1 - \frac{1}{2} c_j \beta_j) \mathbf{E}[(p_s^+)^2].$$

Equation (21) implies that  $p_s^+$  can be expressed as

$$p_s^+ = \frac{1}{\beta^F} (L_s - W_s - q^I)^+ = \frac{1}{\beta^F} (L_s - W_s - q^{I*}(\beta^I, \beta^F))^+.$$

We define  $\beta_{-j} \stackrel{\text{def}}{=} \beta^F - \beta_j$  and write generator  $j$ 's expected profit as

$$\pi_j(\beta_j; \beta_{-j}, \beta^I) = \frac{\beta_j (1 - \frac{1}{2} c_j \beta_j)}{(\beta_j + \beta_{-j})^2} \mathbf{E} \left[ ((L - W - q^{I*}(\beta^I, \beta_j + \beta_{-j}))^+)^2 \right], \quad (\text{A.9})$$

which proves (26).

**Monotonicity of  $q^{I*}(\beta^I, \beta^F)$ :** Recall that the average price  $\bar{P}(q^I, \beta^F)$  decreases in  $q^I$ , as discussed after the definition in (17). Furthermore, because  $L$  and  $W$  have continuous distributions,  $\bar{P}(q^I, \beta^F)$  is differentiable in  $q^I$  everywhere. Denote  $\bar{P}_1 \equiv \partial \bar{P} / \partial q^I$ . We have  $\bar{P}_1 \leq 0$ .

Equation (23),  $q^{I*} - \beta^I \bar{P}(q^{I*}, \beta^F) = 0$ , implicitly determines  $q^{I*}$  as a function of  $\beta_k$ ,  $k \in G^I \cup G^F$ . Therefore,

$$\frac{\partial q^{I*}}{\partial \beta_i} = \frac{\bar{P}}{1 - \beta^I \bar{P}_1} = \frac{q^{I*}}{\beta^I (1 - \beta^I \bar{P}_1)} > 0, \quad i \in G^I, \quad (\text{A.10})$$

which proves that  $q^{I*}$  strictly increases in  $\beta_i$ .

We finally derive  $\partial q^{I*} / \partial \beta_j$ . To simplify notations, let  $D = L - W$  denote the net demand. For a continuous random variable  $X$ , we use  $f_X(x)$  and  $F_X(x)$  to denote the probability density and cumulative distribution functions, respectively, and let  $\bar{F}_X(x) = 1 - F_X(x)$ .

Then, we can write the average price function in (22) as

$$\begin{aligned} \bar{P}(q^I, \beta^F) &= \frac{1}{\beta^F} \mathbf{E}[(D - q^I)^+] - c_h \int_{q^I - \mu(r)}^{q^I} (q^I - x) f_D(x) dx - c_h \int_{-\infty}^{q^I - \mu(r)} (q^I - x) f_L(x) dx \\ &\quad - a_h F_D(q^I) + (a_h - r) F_D(q^I - \mu(r)) + (r - a_h) F_L(q^I - \mu(r)). \end{aligned} \quad (\text{A.11})$$

Thus,

$$\begin{aligned} \bar{P}_2 &\equiv \frac{\partial \bar{P}}{\partial \beta^F} = -\frac{1}{(\beta^F)^2} \mathbf{E}[(D - q^I)^+] < 0, \\ \frac{\partial q^{I*}}{\partial \beta_j} &= \frac{\beta^I \bar{P}_2}{1 - \beta^I \bar{P}_1} = -\frac{\beta^I \mathbf{E}[(D - q^{I*})^+]}{(\beta^F)^2 (1 - \beta^I \bar{P}_1)} < 0, \quad j \in G^F. \end{aligned} \quad (\text{A.12})$$

This completes the proof. ■

**Proof of Corollary 1.** When  $D = L - W > q^{I*}$  (i.e., event  $A_4$ ) occurs with probability one, equations (21)-(22) imply that  $\bar{P}(q^{I*}, \beta^F) = (\mathbf{E}[D] - q^{I*}) / \beta^F$ . Solving equation (23) for  $q^{I*}$  gives

$$q^{I*}(\beta^I, \beta^F) = \frac{\beta^I}{\beta^I + \beta^F} \mathbf{E}[D].$$

Consequently, (25) becomes

$$\pi_i = \frac{\beta_i(1 - \frac{1}{2}c_i\beta_i)}{(\beta^I)^2} \left( \frac{\beta^I}{\beta^I + \beta^F} \mathbb{E}[D] \right)^2 = \frac{\beta_i(1 - \frac{1}{2}c_i\beta_i)}{(\beta^I + \beta^F)^2} (\mathbb{E}[D])^2 = \lambda(i, G^I \cup G^F) (\mathbb{E}[D])^2,$$

and (26) simplifies to

$$\begin{aligned} \pi_j &= \frac{\beta_j(1 - \frac{1}{2}c_j\beta_j)}{(\beta^F)^2} \mathbb{E} \left[ \left( D - \frac{\beta^I}{\beta^I + \beta^F} \mathbb{E}[D] \right)^2 \right] \\ &= \frac{\beta_j(1 - \frac{1}{2}c_j\beta_j)}{(\beta^F)^2} \mathbb{E} \left[ \left( D - \mathbb{E}[D] + \frac{\beta^F}{\beta^I + \beta^F} \mathbb{E}[D] \right)^2 \right] \\ &= \frac{\beta_j(1 - \frac{1}{2}c_j\beta_j)}{(\beta^F)^2} \left( \text{Var}[D] + \left( \frac{\beta^F}{\beta^I + \beta^F} \mathbb{E}[D] \right)^2 \right) \\ &= \lambda(j, G^F) \text{Var}[D] + \lambda(j, G^I \cup G^F) (\mathbb{E}[D])^2. \end{aligned}$$

This completes the proof of Corollary 1. ■

**Proof of Lemma 3.** We first bound the average price in (A.11). Note that

$$\begin{aligned} \int_{q^I - \mu(r)}^{q^I} (q^I - x) f_D(x) dx &\geq 0, \quad \text{and} \\ \int_{-\infty}^{q^I - \mu(r)} (q^I - x) f_L(x) dx &> \mu(r) F_L(q^I - \mu(r)) \\ &= \frac{(r - a_h)^+}{c_h} F_L(q^I - \mu(r)) \geq \frac{(r - a_h)}{c_h} F_L(q^I - \mu(r)). \end{aligned}$$

Using these inequalities, the average price in (A.11) is bounded above by

$$\bar{P}(q^I, \beta^F) < \frac{1}{\beta^F} \mathbb{E} [(D - q^I)^+] - a_h F_D(q^I) + (a_h - r) F_D(q^I - \mu(r)). \quad (\text{A.13})$$

If  $a_h \geq r$ , then  $\mu(r) = 0$  and (A.13) becomes  $\bar{P}(q^I, \beta^F) < \frac{1}{\beta^F} \mathbb{E} [(D - q^I)^+] - r F_D(q^I)$ . If  $a_h < r$ , then (A.13) implies  $\bar{P}(q^I, \beta^F) < \frac{1}{\beta^F} \mathbb{E} [(D - q^I)^+] - a_h F_D(q^I)$ . Combining these two cases, we obtain

$$\bar{P}(q^I, \beta^F) < \frac{1}{\beta^F} \mathbb{E} [(D - q^I)^+] - \min\{r, a_h\} F_D(q^I). \quad (\text{A.14})$$

We can express and bound  $\mathbb{E} [(D - q^I)^+]$  as follows:

$$\begin{aligned} \mathbb{E} [(D - q^I)^+] &= \int_{q^I}^{\infty} (x - \mu_D + \mu_D - q^I) f_D(x) dx \\ &= \int_{q^I}^{\infty} \frac{x - \mu_D}{\sqrt{2\pi}\sigma_D} \exp\left(-\frac{(x - \mu_D)^2}{2\sigma_D^2}\right) dx + \int_{q^I}^{\infty} (\mu_D - q^I) f_D(x) dx \\ &= \frac{\sigma_D}{\sqrt{2\pi}} \int_{\frac{q^I - \mu_D}{\sigma_D}}^{\infty} y \exp(-y^2/2) dy + (\mu_D - q^I) \bar{F}_D(q^I) \\ &\leq \frac{\sigma_D}{\sqrt{2\pi}} + (\mu_D - q^I) \bar{F}_D(q^I). \end{aligned} \quad (\text{A.15})$$

The inequalities (A.14) and (A.15) lead to

$$\bar{P}(q^I, \beta^F) < \frac{\sigma_D}{\sqrt{2\pi}\beta^F} + \frac{(\mu_D - q^I) \bar{F}_D(q^I)}{\beta^F} - \min\{r, a_h\} F_D(q^I). \quad (\text{A.16})$$

Using (23) and (A.16), we have

$$\begin{aligned} q^{I \max} &= \beta^{I \max} \overline{P}(q^{I \max}, \beta^{F \min}) \\ &< \beta^{I \max} \left[ \frac{\sigma_D}{\sqrt{2\pi}\beta^{F \min}} + \frac{(\mu_D - q^{I \max})\overline{F}_D(q^{I \max})}{\beta^{F \min}} - \min\{r, a_h\}F_D(q^{I \max}) \right]. \end{aligned} \quad (\text{A.17})$$

We now prove  $q^{I \max} < \mu_D$ . If the opposite is true,  $q^{I \max} \geq \mu_D$ , then  $F_D(q^{I \max}) \geq \frac{1}{2}$  and (A.17) implies

$$q^{I \max} < \beta^{I \max} \left[ \frac{\sigma_D}{\sqrt{2\pi}\beta^{F \min}} - \frac{\min\{r, a_h\}}{2} \right] \leq \beta^{I \max} \left[ \frac{\sigma_D^*}{\sqrt{2\pi}\beta^{F \min}} - \frac{\min\{r, a_h\}}{2} \right] = \mu_D,$$

where  $\sigma_D^* \equiv \sqrt{2\pi}\beta^{F \min} \left[ \frac{\mu_D}{\beta^{I \max}} + \frac{\min\{r, a_h\}}{2} \right]$ . This contradicts  $q^{I \max} \geq \mu_D$ . Therefore, we conclude that  $q^{I \max} < \mu_D$  when  $\sigma_D \leq \sigma_D^*$ .  $\blacksquare$

**Proof of Theorem 4.** Because generator  $k$ 's pure strategy space is a finite interval  $[\beta_k^{\min}, c_k^{-1}]$ , it suffices to show that,  $\forall k \in G^I \cup G^F$ , generator  $k$ 's profit function is quasi-concave with respect to  $\beta_k$  to prove the existence of a pure strategy Nash equilibrium (Debreu 1952).

The proof of the quasi-concavity will use the derivatives of  $\overline{P}(q^I, \beta^F)$ . Differentiating  $\overline{P}(q^I, \beta^F)$  in (A.11) with respect to  $q^I$  and using  $\mu(r) = (r - a_h)^+/c_h$ , we obtain

$$\begin{aligned} \overline{P}_1(q^I, \beta^F) &\equiv \frac{\partial \overline{P}}{\partial q^I} = -\frac{1}{\beta^F} \overline{F}_D(q^I) - c_h [F_D(q^I) - F_D(q^I - \mu(r)) + F_L(q^I - \mu(r))] \\ &\quad - a_h f_D(q^I) + [c_h \mu(r) + (a_h - r)] [f_D(q^I - \mu(r)) - f_L(q^I - \mu(r))] \\ &= -\frac{1}{\beta^F} \overline{F}_D(q^I) - c_h [F_D(q^I) - F_D(q^I - \mu(r)) + F_L(q^I - \mu(r))] \\ &\quad - a_h f_D(q^I) + (a_h - r)^+ [f_D(q^I - \mu(r)) - f_L(q^I - \mu(r))], \end{aligned} \quad (\text{A.18})$$

$$\begin{aligned} \overline{P}_{11}(q^I, \beta^F) &\equiv \frac{\partial^2 \overline{P}}{\partial q^{I2}} = \frac{1}{\beta^F} f_D(q^I) - c_h [f_D(q^I) - f_D(q^I - \mu(r)) + f_L(q^I - \mu(r))] \\ &\quad - a_h f'_D(q^I) + (a_h - r)^+ [f'_D(q^I - \mu(r)) - f'_L(q^I - \mu(r))]. \end{aligned} \quad (\text{A.19})$$

By Lemma 3, if  $\sigma_D \leq \sigma_D^*$ , then we have  $q^{I \max} < \mu_D$ . When  $q^I < \mu_D$ , and  $\sigma_D \rightarrow 0$ , all the distribution functions in (A.18)-(A.19) approach zero, except for  $\overline{F}_D(q^I)$ , which approaches one. Therefore, when  $\sigma_D$  is small,  $\overline{P}_1$  is close to  $-1/\beta^F$  and  $\overline{P}_{11}$  is close to zero.

**Quasi-concavity of IG's profit function.** The profit function of IG  $i \in G^I$  is expressed as  $\pi_i(\beta_i; \beta_{-i}, \beta^F)$  in (25). To prove its quasi-concavity in  $\beta_i$ , we will show that its derivative  $\partial \pi_i / \partial \beta_i$  can cross zero value from above at most once as  $\beta_i$  increases, while holding  $\beta_{-i}$  and  $\beta^F$  constant.

Noting  $\partial q^{I*} / \partial \beta_i$  in (A.10), we differentiate (25) with respect to  $\beta_i$  to obtain

$$\begin{aligned} \frac{\partial \pi_i}{\partial \beta_i} &= \frac{\beta_i(2 - c_i\beta_i)}{(\beta_i + \beta_{-i})^2} q^{I*} \frac{\partial q^{I*}}{\partial \beta_i} + \frac{\beta_{-i}(1 - c_i\beta_i) - \beta_i}{(\beta_i + \beta_{-i})^3} (q^{I*})^2 \\ &= \frac{\beta_i(2 - c_i\beta_i)}{(\beta^I)^2} \frac{(q^{I*})^2}{\beta^I(1 - \beta^I \overline{P}_1)} + \frac{\beta_{-i}(1 - c_i\beta_i) - \beta_i}{(\beta^I)^3} (q^{I*})^2 \end{aligned}$$

$$\begin{aligned}
&= \frac{(q^{I*})^2}{(\beta^I)^3(1 - \beta^I \bar{P}_1)} \left[ \beta_i(2 - c_i \beta_i) + (\beta_{-i}(1 - c_i \beta_i) - \beta_i)(1 - \beta^I \bar{P}_1) \right] \\
&= \frac{(q^{I*})^2}{(\beta^I)^3(1 - \beta^I \bar{P}_1)} \left[ \beta^I(1 - c_i \beta_i) - (\beta_{-i}(1 - c_i \beta_i) - \beta_i) \beta^I \bar{P}_1 \right] \\
&= \frac{(q^{I*})^2}{(\beta^I)^2(1 - \beta^I \bar{P}_1)} X(\beta_i; \beta_{-i}, \beta^F),
\end{aligned}$$

where  $X(\beta_i; \beta_{-i}, \beta^F) \stackrel{\text{def}}{=} 1 - c_i \beta_i + (\beta_i(1 + c_i \beta_{-i}) - \beta_{-i}) \bar{P}_1$ . To show  $\partial \pi_i / \partial \beta_i$  can cross zero value from above at most once, it suffices to show  $X$  decreases in  $\beta_i$ . Differentiating  $X$  with respect to  $\beta_i$ , we have

$$\frac{\partial X}{\partial \beta_i} = -c_i + (1 + c_i \beta_{-i}) \bar{P}_1 + (\beta_i(1 + c_i \beta_{-i}) - \beta_{-i}) \bar{P}_{11} \frac{q^{I*}}{\beta^I(1 - \beta^I \bar{P}_1)},$$

where  $\bar{P}_{11}$  is derived in (A.19). Note that  $-c_i + (1 + c_i \beta_{-i}) \bar{P}_1 < 0$ . Thus, if  $\bar{P}_{11}(q^{I*}, \beta^F)$  is sufficiently small, we can establish  $\partial X / \partial \beta_i \leq 0$ . Based on the discussion after (A.18) and (A.19), there exists  $\hat{\sigma}_D$ , such that when  $\sigma_D < \hat{\sigma}_D$ , we have  $\partial X / \partial \beta_i \leq 0$  and, therefore,  $\pi_i$  is quasi-concave in  $\beta_i$ .

**Quasi-concavity of FG's profit function.** Denote  $D = L - W$  and  $q^{I*} = q^{I*}(\beta^I, \beta^F)$ , we can write FG's profit function in (26) as

$$\pi_j(\beta_j; \beta_{-j}, \beta^I) = \frac{\beta_j(1 - \frac{1}{2}c_j \beta_j)}{(\beta_j + \beta_{-j})^2} \mathbf{E} \left[ ((D - q^{I*})^+)^2 \right].$$

We will show that  $\partial \pi_j / \partial \beta_j$  can cross zero value at most once from above when  $\beta_j$  increases.

Differentiating  $\pi_j$  with respect to  $\beta_j$  and using  $\partial q^{I*} / \partial \beta_j$  from (A.12) and the following fact

$$\begin{aligned}
\frac{\partial}{\partial q^{I*}} \mathbf{E} \left[ ((D - q^{I*})^+)^2 \right] &= \frac{\partial}{\partial q^{I*}} \int_{q^{I*}}^{\infty} (x - q^{I*})^2 f_D(x) dx \\
&= \int_{q^{I*}}^{\infty} -2(x - q^{I*}) f_D(x) dx = -2 \mathbf{E}[(D - q^{I*})^+],
\end{aligned}$$

we obtain

$$\begin{aligned}
\frac{\partial \pi_j}{\partial \beta_j} &= \frac{\beta_{-j}(1 - c_j \beta_j) - \beta_j}{(\beta^F)^3} \mathbf{E} \left[ ((D - q^{I*})^+)^2 \right] + \frac{\beta_j(2 - c_j \beta_j)}{(\beta^F)^2} \mathbf{E}[(D - q^{I*})^+] \frac{\beta^I \mathbf{E}[(D - q^{I*})^+]}{(\beta^F)^2(1 - \beta^I \bar{P}_1)} \\
&= \frac{\beta_j \mathbf{E} \left[ ((D - q^{I*})^+)^2 \right]}{(\beta^F)^3(1 - \beta^I \bar{P}_1)} \left[ Y(\beta_j, \beta_{-j}, \beta^I) + \beta^I Z(\beta_j, \beta_{-j}, \beta^I) \right],
\end{aligned}$$

where

$$\begin{aligned}
Y(\beta_j, \beta_{-j}, \beta^I) &\stackrel{\text{def}}{=} \left( \frac{\beta_{-j}}{\beta_j} - (1 + c_j \beta_{-j}) \right) (1 - \beta^I \bar{P}_1), \\
Z(\beta_j, \beta_{-j}, \beta^I) &\stackrel{\text{def}}{=} \frac{2 - c_j \beta_j}{\beta_j + \beta_{-j}} \psi(q^{I*}), \\
\psi(q^{I*}) &\stackrel{\text{def}}{=} \frac{(\mathbf{E}[(D - q^{I*})^+])^2}{\mathbf{E}[(D - q^{I*})^+)^2].}
\end{aligned}$$

It suffices to show that  $Y$  and  $Z$  decrease in  $\beta_j$ . Differentiating  $Y$  with respect to  $\beta_j$ , we have

$$\frac{\partial Y}{\partial \beta_j} = -\frac{\beta_{-j}}{\beta_j^2} (1 - \beta^I \bar{P}_1) + \left( \frac{\beta_{-j}}{\beta_j} - (1 + c_j \beta_{-j}) \right) \beta^I \bar{P}_{11} \frac{\beta^I \mathbf{E}[(D - q^{I*})^+]}{(\beta^F)^2(1 - \beta^I \bar{P}_1)}.$$

By the same argument used for the quasi-concavity of  $\pi_i$ , we see that when  $\sigma_D$  is sufficiently small,  $\bar{P}_1$  is close to  $-1/\beta^F$  and  $\bar{P}_{11}$  is close to zero. Thus, there exists  $\tilde{\sigma}_D$ , such that when  $\sigma_D < \tilde{\sigma}_D$ , we have  $\partial Y/\partial\beta_j \leq 0$ .

Next, we show that  $Z$  decreases in  $\beta_j$ . Note that  $\frac{\partial}{\partial q^{I*}}\mathbb{E}[(D - q^{I*})^+] = -\bar{F}_D(q^{I*})$  and

$$\psi'(q^{I*}) = -\frac{2\mathbb{E}[(D - q^{I*})^+]\bar{F}_D(q^{I*})}{\mathbb{E}[(D - q^{I*})^+]^2} + \frac{2(\mathbb{E}[(D - q^{I*})^+])^3}{(\mathbb{E}[(D - q^{I*})^+]^2)^2} = -\frac{2\psi(q^{I*})(\bar{F}_D(q^{I*}) - \psi(q^{I*}))}{\mathbb{E}[(D - q^{I*})^+]}$$

Using this derivative and  $\partial q^{I*}/\partial\beta_j$  in (A.12), we have

$$\begin{aligned} \frac{\partial Z}{\partial\beta_j} &= \frac{-c_j\beta^F - (2 - c_j\beta_j)}{(\beta^F)^2} \psi(q^{I*}) + \frac{2 - c_j\beta_j}{\beta^F} \psi'(q^{I*}) \frac{\partial q^{I*}}{\partial\beta_j} \\ &= \frac{-c_j\beta^F - (2 - c_j\beta_j)}{(\beta^F)^2} \psi(q^{I*}) + \frac{2 - c_j\beta_j}{\beta^F} \frac{2\psi(q^{I*})(\bar{F}_D(q^{I*}) - \psi(q^{I*}))\beta^I}{(\beta^F)^2(1 - \beta^I\bar{P}_1)} \\ &= \frac{\psi(q^{I*})}{(\beta^F)^2} \left[ -c_j\beta^F - (2 - c_j\beta_j) \left( 1 - \frac{2(\bar{F}_D(q^{I*}) - \psi(q^{I*}))\beta^I}{\beta^F(1 - \beta^I\bar{P}_1)} \right) \right]. \end{aligned}$$

We will show that  $\psi(q^{I*})$  is close to  $\bar{F}_D(q^{I*})$  when  $\sigma_D$  is sufficiently small and  $q^{I*} < \mu_D$  to complete the proof.

For a normal random variable  $X \sim \mathcal{N}(\mu, \sigma)$ , we can show that  $\mathbb{E}[X^+] = \mu\bar{F}_X(0) + \sigma^2 f_X(0)$  and  $\mathbb{E}[(X^+)^2] = (\mu^2 + \sigma^2)\bar{F}_X(0) + \mu\sigma^2 f_X(0)$ . Then,

$$\begin{aligned} \mathbb{E}[(D - q^{I*})^+] &= (\mu_D - q^{I*})\bar{F}_D(q^{I*}) + \sigma_D^2 f_D(q^{I*}), \\ \mathbb{E}[(D - q^{I*})^+]^2 &= ((\mu_D - q^{I*})^2 + \sigma_D^2)\bar{F}_D(q^{I*}) + \sigma_D^2(\mu_D - q^{I*})f_D(q^{I*}), \\ \psi(q^{I*}) &= \frac{[(\mu_D - q^{I*})\bar{F}_D(q^{I*}) + \sigma_D^2 f_D(q^{I*})]^2}{[(\mu_D - q^{I*})^2 + \sigma_D^2]\bar{F}_D(q^{I*}) + \sigma_D^2(\mu_D - q^{I*})f_D(q^{I*})}. \end{aligned}$$

If  $\sigma_D \leq \sigma_D^*$ , we have  $q^{I\max} < \mu_D$  (Lemma 3). The above expression for  $\psi(q^{I*})$  implies that as  $\sigma_D \rightarrow 0$ , we have  $\bar{F}_D(q^{I*}) \rightarrow 1$ ,  $f_D(q^{I*}) \rightarrow 0$ , and  $\psi(q^{I*}) \rightarrow 1$ . Hence, there exists  $\sigma_D^\dagger$ , such that when  $\sigma_D < \sigma_D^\dagger$ , we have  $\partial Z/\partial\beta_j \leq 0$ .

To summarize, when  $\sigma_D < \min\{\sigma_D^*, \hat{\sigma}_D, \tilde{\sigma}_D, \sigma_D^\dagger\}$ , the profit function  $\pi_j$  is quasi-concave in  $\beta_j$ . This establishes the existence of a pure strategy equilibrium, i.e., the linear supply function equilibrium. ■

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