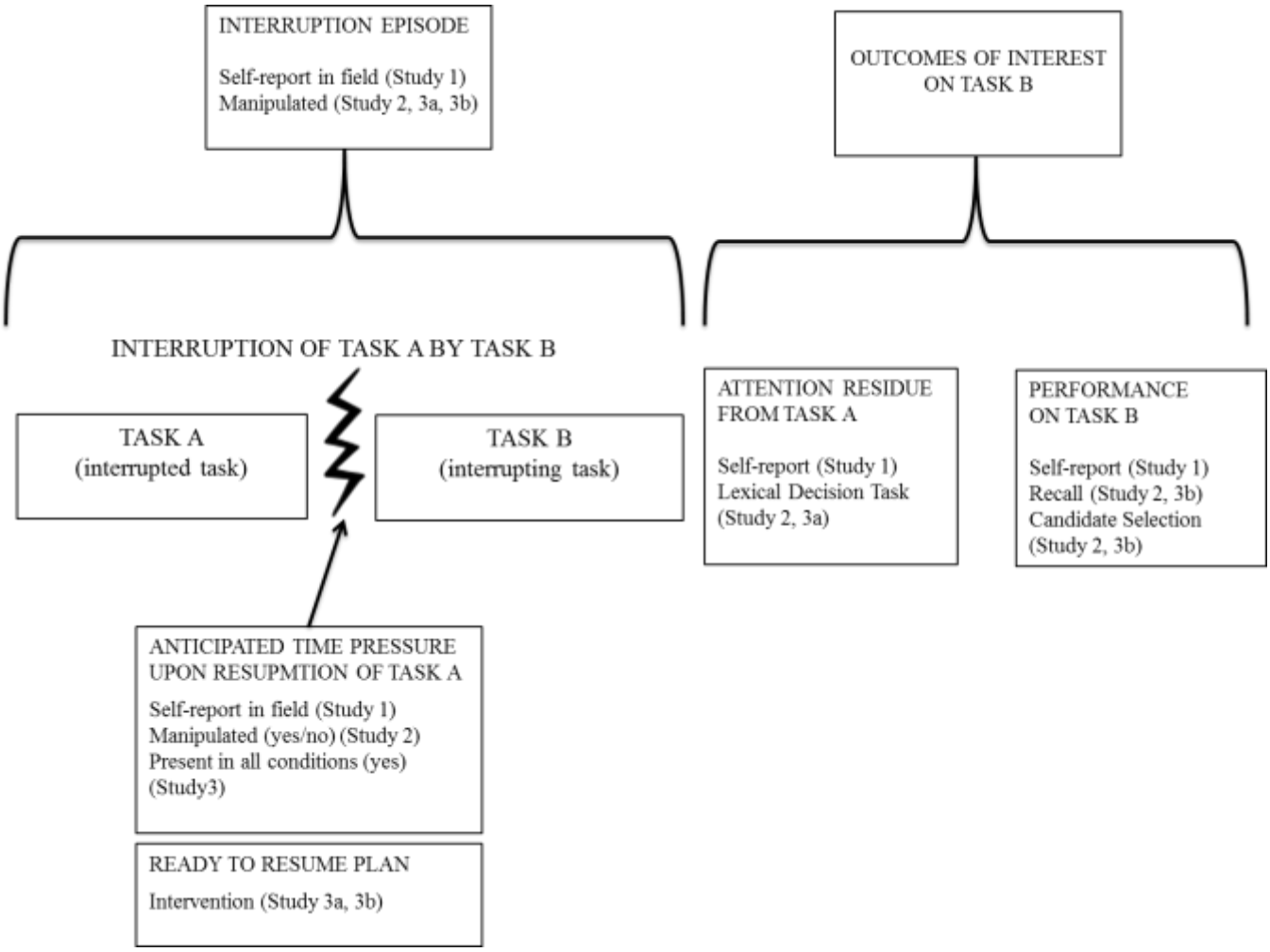
TASK INTERRUPTED: HOW ANTICIPATING TIME PRESSURE ON RESUMPTION OF AN INTERRUPTED TASK CAUSES ATTENTION RESIDUE AND LOW PERFORMANCE ON INTERRUPTING TASKS AND HOW A "READY-TO-RESUME" PLAN MITIGATES THE EFFECTS

Online Appendix A – Schematic representation of our study designs.



Online Appendix B – Study 1 Supplemental Analysis

In Study 1, common method bias represents an important limitation and a compelling alternative explanation. Anticipating this issues, we used procedural and statistical remedies from Podsakoff et al. 2003. First, from a procedural perspective, we designed Study 1 to follow their recommendations about assuring survey anonymity and separating predictor and criterion measurements, along with recommendations of Beal et al. (2005) and Fowler (1995) about assessing performance. Specifically, surveys were anonymous and collected outside work with no relation to the employer, reducing incentives to understate or embellish responses (decreased evaluation apprehension). Furthermore, we did not ask measures in model sequence: instead we asked about attention residue and performance measures prior to task characteristics to separate predictor and criterion measurement. Most of the variables regarding interrupted and interrupting tasks were factual items, not attitudinal or perceptual, and we gathered self-reported performance relative to personal normative performance levels. The focus on factual task characteristics was also important to minimize potential retrospection issues. Furthermore, by focusing on the last interruption of the day, we minimized time lag between event and response. We also asked respondents to first briefly describe the tasks before reporting on task characteristics and outcome variables to facilitate recall. The specificity, detail, and factual nature of their descriptions add confidence in the quality of the data. Thus, design features and scale item quality diminished common method variance and recall concerns.

Second, we performed empirical tests to verify whether our data possessed common variance bias. We conducted Harman's single-factor test using indicators from our main variables of interest. Our unrotated principal component analysis showed no general factor, suggesting common method bias did not seem to be a problem. The first factor accounted for 21% and did not account for the majority of the variance. However, Podsakoff et al. (2003) suggested that Harman's method is a weak detection tool. Therefore, we followed up with the single common method factor approach to test and compare the fit of (1) Model 1, a base model with all items loading on their respective factors; (2) Model 2, the base model

with the addition of an unmeasured method factor on which all items were also set to load; and (3) Model 3, which was the same as Model 2 with the addition of a marker that lacked correlation with our variables of interest and whose item was set to load on the common method factor. The three models fit the data well (CFI=.94 and RMSEA=.05). However, comparing Models 1 and 2, the method factor did not significantly improve model fit and the common factor variance was only 10%, which is much less than the 25% reported by Williams et al. (1989) based on analysis across 11 data sets. The results suggest that common method variance is not a serious issue here. Adding a marker reduced the variance to 9% (Model 3). Reduction in variance further confirms that common method variance is not a serious concern.